

The technical argument in the final paragraph of this proof can be better appreciated after studying Chapter 18. It is neither possible nor necessary in this approach to assert that  $E(W_m|\mathcal{F}_{t-1}) \rightarrow 0$  a.s.

Note how taking conditional expectations of (16.12) yields

$$E(X_t|\mathcal{F}_{t-1}) = Z_t - Z_{t+1} \text{ a.s.} \quad (16.21)$$

It follows that  $W_t$  is almost surely equal to the centred r.v.  $X_t - E(X_t|\mathcal{F}_{t-1})$ .

**16.7 Example** Consider the linear process from **16.2**, with  $\{U_t\}$  a stationary integrable sequence. Then  $X_t$  is stationary, and

$$E|X_1| \leq E|U_1| \sum_{j=-\infty}^{\infty} |\theta_j| < \infty.$$

If the coefficients satisfy a stronger summability condition, i.e.

$$\sum_{m=1}^{\infty} \sum_{j=m}^{\infty} (|\theta_j| + |\theta_{-j}|) = \sum_{m=1}^{\infty} m|\theta_m| + \sum_{m=1}^{\infty} m|\theta_{-m}| < \infty, \quad (16.22)$$

then  $X_t$  is an  $L_1$ -mixingale of size  $-1$ . By a rearrangement of terms we obtain the decomposition of (16.12) with

$$W_t = \left( \sum_{j=-\infty}^{\infty} \theta_j \right) U_t \quad (16.23)$$

and

$$Z_t = \sum_{m=1}^{\infty} \left( \left( \sum_{j=m}^{\infty} \theta_j \right) U_{t-m} - \left( \sum_{j=m}^{\infty} \theta_{-j} \right) U_{t+m-1} \right), \quad (16.24)$$

where  $E|Z_t| < \infty$  by (16.22).  $\square$

### 16.3 Maximal Inequalities

As with martingales, maximal inequalities are central to applications of the mixingale concept in limit theory. The basic idea of these results is to extend Doob's inequality (**15.15**) by exploiting the representation as a telescoping sum of martingale differences. McLeish's idea is to let  $m$  go to  $\infty$  in (16.10), and accordingly write

$$S_n = \sum_{k=-\infty}^{\infty} Y_{nk}, \text{ a.s.} \quad (16.25)$$

**16.8 Lemma** Suppose  $\{S_j\}_1^n$  has the representation in (16.25). Let  $\{a_k\}_{-\infty}^{\infty}$  be a summable collection of non-negative real numbers, with  $a_k = 0$  if  $Y_{nk} = 0$  a.s., and  $a_k > 0$  otherwise. For any  $p > 1$ ,

$$E \left( \max_{1 \leq j \leq n} |S_j|^p \right) \leq \left( \frac{p}{p-1} \right)^p \left( \sum_{k=-\infty}^{\infty} a_k \right)^{p-1} \sum_{a_k > 0} a_k^{1-p} E|Y_{nk}|^p. \quad (16.26)$$